



Transition Models

for Precipitation Climatology Estimation

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A Journey in Probabilistic Modeling

- Over the past 15 years, we have developed efficient algorithms and open-source software for (Bayesian) distributional regression, culminating in the CRAN package *bamlss*.
- Distributional regression provides a highly flexible framework for full probabilistic modeling of complex data.
- Applications span a wide range: univariate and multivariate responses, count data, censored and survival outcomes, joint models, and more.
- At UIBK, long-term research projects focus on count data modeling, e.g., for probabilistic forecasting of lightning strike counts.

Model specification

Any parameter of a population distribution $\mathcal{D}$ may be modeled by explanatory variables

$$y \sim \mathcal{D}(\theta_1(\mathbf{x}; \beta_1), \dots, \theta_K(\mathbf{x}; \beta_K)),$$



with $\beta = (\beta_1^\top, \dots, \beta_K^\top)^\top$.

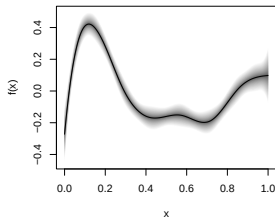
Each parameter is linked to a structured additive predictor

$$h_k(\theta_k(\mathbf{x}; \beta_k)) = f_{1k}(\mathbf{x}; \beta_{1k}) + \dots + f_{J_k k}(\mathbf{x}; \beta_{J_k k}); \quad j = 1, \dots, J_k; \quad k = 1, \dots, K.$$

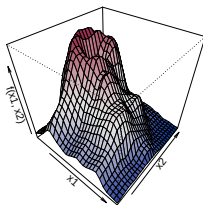
- $h_k(\cdot)$: Link functions for each distribution parameter.
- $f_{jk}(\cdot)$: Model terms of one or more variables.

Model Terms $f_{jk}(\cdot)$

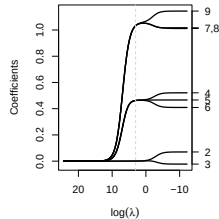
Nonlinear Effects



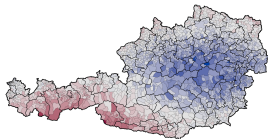
Two-Dimensional Surfaces



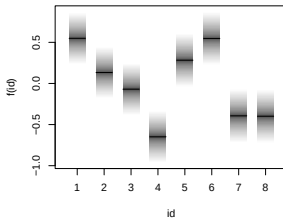
LASSO & Factor Clustering



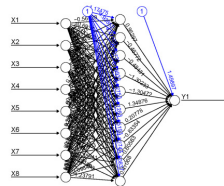
Spatially Correlated Effects $f(x) = f(s)$



Random Intercepts $f(x) = f(id)$

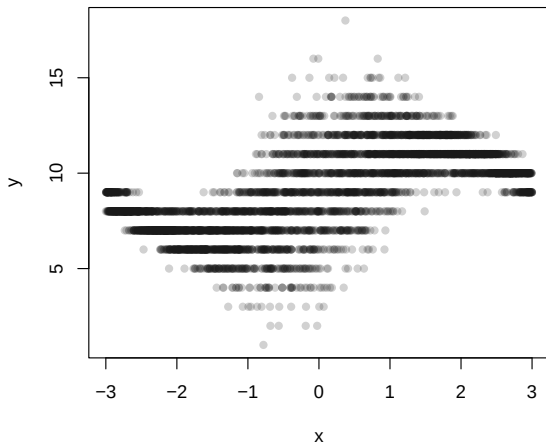


Neural Networks



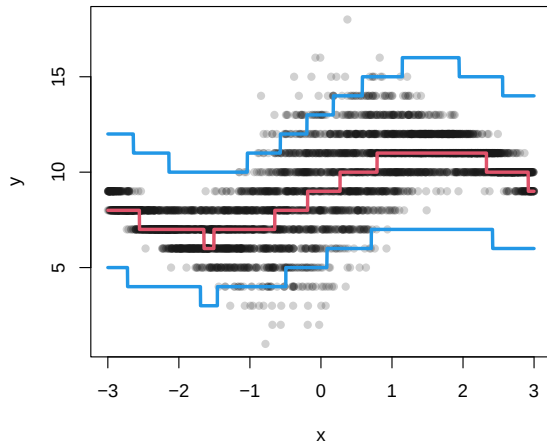
Count Models

Simulated example: $y \sim \text{NO}(\mu = f(x), \log(\sigma) = f(x))$, transformed to counts.



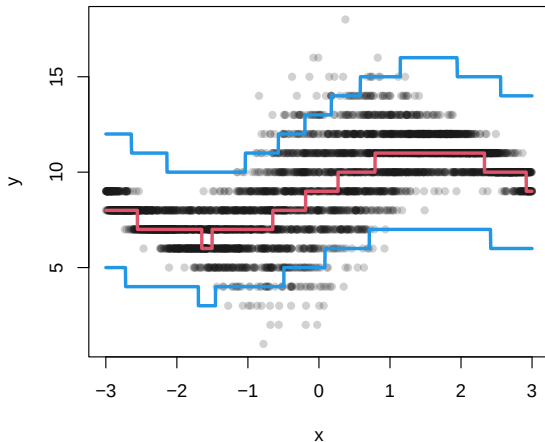
Count Models

Estimated model: $y \sim \text{PO}(\log(\mu) = f(x))$.



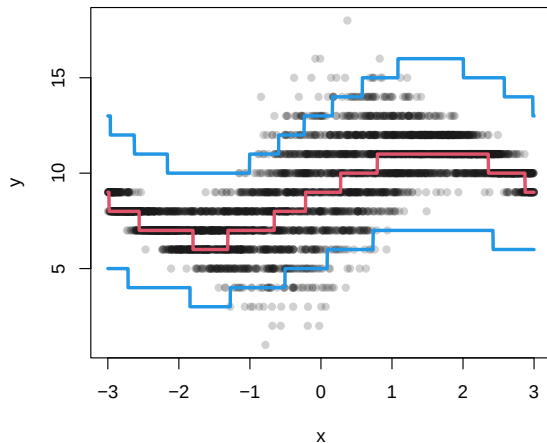
Count Models

Estimated model: $y \sim \text{NBI}(\log(\mu) = f(x), \log(\sigma) = f(x))$.



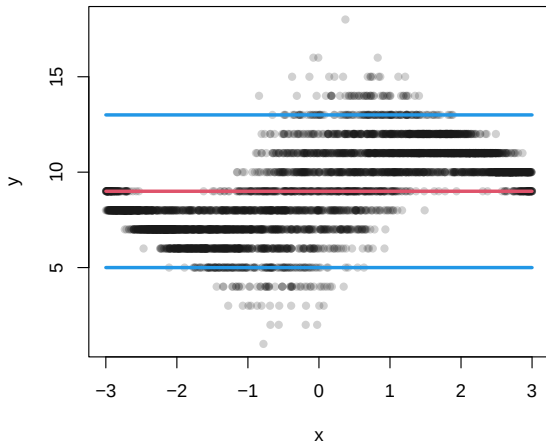
Count Models

Estimated model: $y \sim \text{SICHEL}(\log(\mu) = f(x), \log(\sigma) = f(x), \nu = f(x)).$



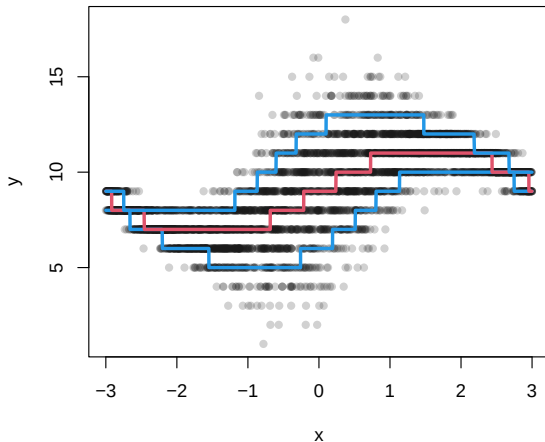
Count Models

Estimated model: $y \sim \text{GPO}(\log(\mu) = f(x), \log(\sigma) = f(x)).$



Count Models

Estimated model: $y \sim \text{DPO}(\log(\mu) = f(x), \log(\sigma) = f(x))$.



Transition Models

The transition probability $P(\cdot)$ for count data is defined as

$$P(y_i > r | y_i \geq r, \mathbf{x}_i) = F(\eta_{ir}(\alpha)), \quad r = 0, 1, 2, \dots$$

where $F(\cdot)$ is a CDF (e.g., logistic or probit) and r represents the counts, with an additive predictor

$$\eta_{ir}(\alpha) = \theta_r + \sum_{j=1}^k f_j(\mathbf{x}_i, r; \beta).$$

The parameters $\alpha = (\theta^\top, \beta^\top)$ include count-specific intercepts and (possibly) smooth functions $f_j(\cdot)$. For i.i.d. observations, let π_{ir} denote the probability that the count response equals r , i.e., $P(y_i = r | \mathbf{x}_i)$. These probabilities are computed recursively as

$$\pi_{ir} = (1 - F(\eta_{ir}(\alpha))) \prod_{s=0}^{r-1} F(\eta_{is}(\alpha)).$$

Transition Models

Parameter estimation considers the underlying Markov chain $Y_{i0}, Y_{i1}, \dots$, where

$$Y_{ir} = 1 - I(y_i = r).$$

Simplifies to binary model log-likelihood

$$\ell(\boldsymbol{\alpha}) = \sum_{i=1}^n \sum_{s=0}^{y_i} \left[Y_{is} \log(F(\eta_{ir})) + (1 - Y_{is}) \log(1 - F(\eta_{ir})) \right].$$

$(Y_{i0}, \dots, Y_{iy_i})^\top = (1, \dots, 1, 0)$ are created, along with a new covariate $\theta_i = (0, 1, 2, \dots, y_i)^\top$ to capture count-specific effects $f_j(\mathbf{x}_i, \theta_i)$, or simple count-specific intercepts.

All other covariates are duplicated accordingly.

Example

In R:

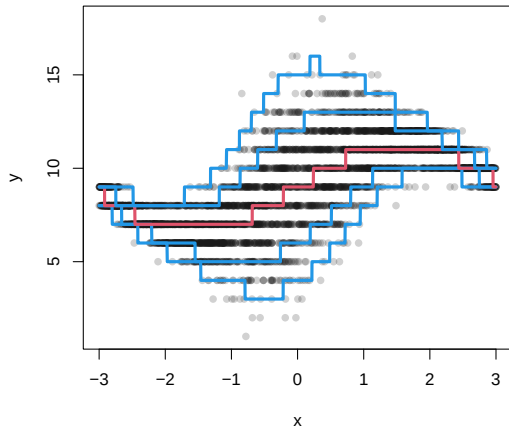
```
R> print(head(df, 10))
```

	index	myresponse	Y	theta	x	y	z
1	1	5	1	0	-1.2	765	foo
2	1	5	1	1	-1.2	765	foo
3	1	5	1	2	-1.2	765	foo
4	1	5	1	3	-1.2	765	foo
5	1	5	1	4	-1.2	765	foo
6	1	5	0	5	-1.2	765	foo
7	2	0	0	0	3.2	731	bar
8	3	2	1	0	-0.5	353	foo
9	3	2	1	1	-0.5	353	foo
10	3	2	0	2	-0.5	353	foo

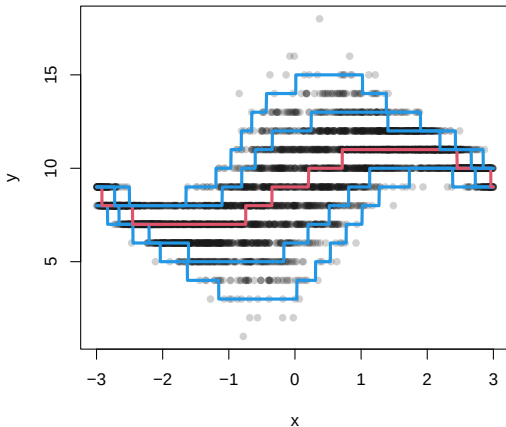
```
R> b <- glm(Y ~ as.factor(theta) + x + y + z, data = df, family = binomial)
```

Count Models

DPO count model

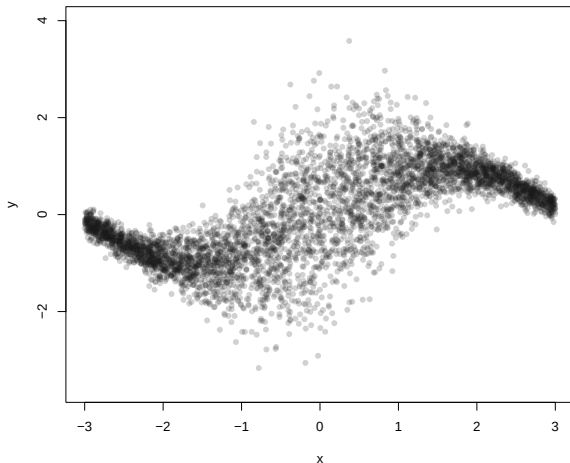


Transition model



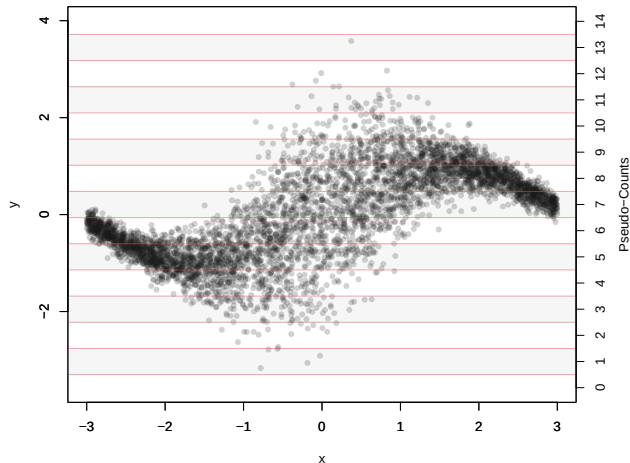
Extension for Continuous Responses

Simulated example: $y \sim \text{NO}(\mu = f(x), \log(\sigma) = f(x))$.



Extension for Continuous Responses

Idea: Define intervals and generate pseudo-counts.



Extension for Continuous Responses

Discretization approach inspired by histogram binning.

Divide response $y_i \in \mathbb{R}$, $i = 1, \dots, n$, into $m - 1$ intervals using

$$\zeta_1, \zeta_2, \dots, \zeta_m,$$

where each interval $[\zeta_r, \zeta_{r+1})$ corresponds to a discrete count r .

Each observation y_i is assigned a pseudo count $\tilde{y}_i$.

For a continuous response variable y_i with CDF $F(y)$, discretization process approximates the probabilities of y_i falling into each interval as

$$P(\zeta_r \leq y_i < \zeta_{r+1}) = F(\zeta_{r+1}) - F(\zeta_r).$$

Extension for Continuous Responses

Probabilities are encoded as transformed counts $\tilde{y}_i$, so that the transition model uses

$$P(\tilde{y}_i = r) = P(\zeta_r \leq y_i < \zeta_{r+1})$$

to approximate the discrete likelihood.

The transition model estimates the probability of transitions between counts

$$P(\tilde{y}_i > r \mid \tilde{y}_i \geq r, \mathbf{x}_i) = F(\eta_{ir}(\boldsymbol{\alpha}))$$

and recursively computes

$$P(\tilde{y}_i = r, \mathbf{x}_i) = P(\tilde{y}_i = r \mid \tilde{y}_i \geq r, \mathbf{x}_i) \prod_{s=0}^{r-1} P(\tilde{y}_i > s \mid \tilde{y}_i \geq s, \mathbf{x}_i).$$

Extension for Continuous Responses

- Let r denote the unique index such that $y_i \in [\zeta_r, \zeta_{r+1})$.
- For any value $y_i \in [\zeta_r, \zeta_{r+1})$, the CDF can be approximated by

$$\hat{F}(y_i) = \sum_{s=0}^{r-1} P(\tilde{y}_i = s) + \frac{y_i - \zeta_r}{\zeta_{r+1} - \zeta_r} P(\tilde{y}_i = r).$$

- The PDF can be approximated as

$$\hat{f}(y_i) = \frac{P(\tilde{y}_i = r)}{\zeta_{r+1} - \zeta_r}.$$

- The mean and variance are approximated using midpoints $c_r = \frac{\zeta_r + \zeta_{r+1}}{2}$

$$E[Y] = \sum_r c_r P(\tilde{y} = r), \quad \text{Var}(Y) = \sum_r c_r^2 P(\tilde{y} = r) - \left(\sum_r c_r P(\tilde{y} = r) \right)^2.$$

Extension for Continuous Responses

- Similarly: Skewness = $\frac{E[(Y-E[Y])^3]}{\text{Var}(Y)^{3/2}}$, Kurtosis = $\frac{E[(Y-E[Y])^4]}{\text{Var}(Y)^2}$.
- The τ -quantile $\hat{Q}(\tau)$ is obtained by finding the smallest index r such that

$$\sum_{s=0}^r P(\tilde{y} = s) \geq \tau,$$

using linear interpolation within the interval

$$\hat{Q}(\tau) = \zeta_r + \frac{\tau - \sum_{s=0}^{r-1} P(\tilde{y} = s)}{P(\tilde{y} = r)} \cdot (\zeta_{r+1} - \zeta_r).$$

- The mode $\hat{M}$ is approximated as

$$\hat{M} = c_{r^*}, \quad \text{where} \quad r^* = \arg \max_r P(\tilde{y} = r).$$

Software

An implementation is provided in the R package *transitreg*.

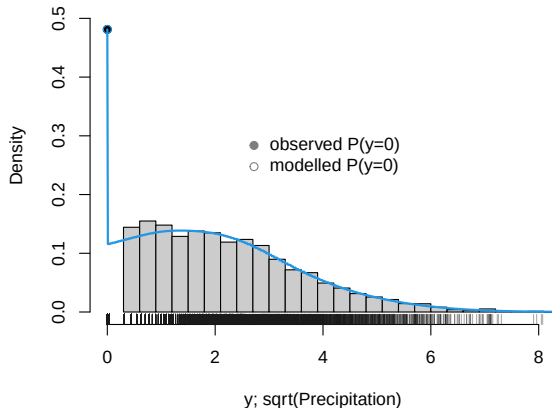
<https://github.com/retostauffer/transitreg>

Install with:

```
R> install.packages("transitreg",  
+   repos = c("https://gamlss-dev.R-universe.dev",  
+             "https://cloud.R-project.org"))
```

Application: Precipitation Climatology Estimation

```
R> b <- transitreg(sqrt_pre ~ s(theta), data = df, breaks = 30, censored = "left")
```



Application: Precipitation Climatology Estimation

Model formula.

```
R> f <- sqrt_pre ~ theta0 + s(theta, k = 20) + s(day, bs = "cc", k = 20) +  
+   te(theta, day, bs = c("cr", "cc"), k = 10)
```

Estimate model.

```
R> breaks <- seq(0, 10, by = 0.05)  
R> b <- transitreg(f, data = dtrain, breaks = breaks, censored = "left")
```

Predict quantiles.

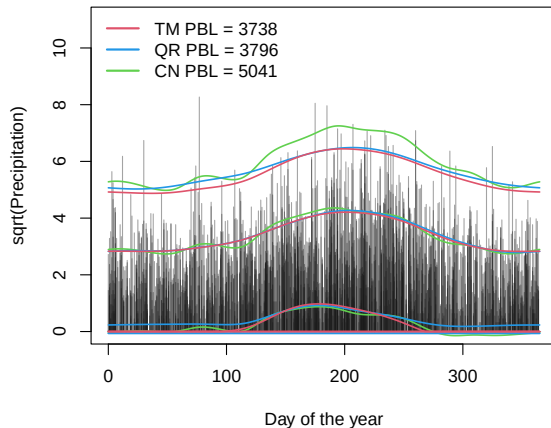
```
R> p <- predict(b, newdata = dtest, prob = c(0.01, 0.1, 0.5, 0.9, 0.99))
```

Compare with *qgam*.

```
R> library("qgam")  
R> qu <- c(0.01, 0.1, 0.5, 0.9, 0.99)  
R> m <- mqgam(sqrt_pre ~ s(day, k = 20, bs = "cc"), data = dtrain, qu = qu)
```

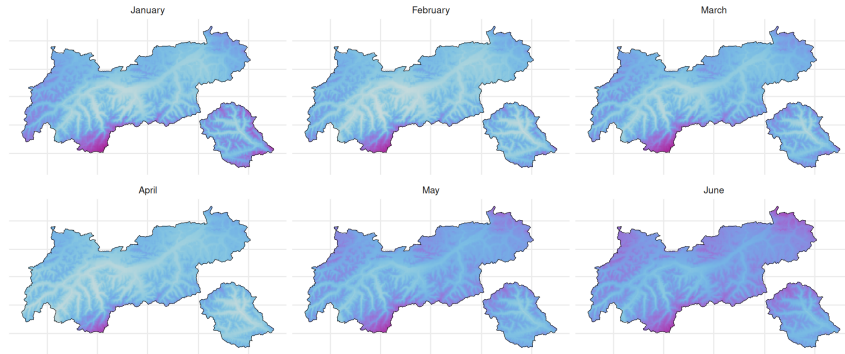
Application: Precipitation Climatology Estimation

Estimated climatology.



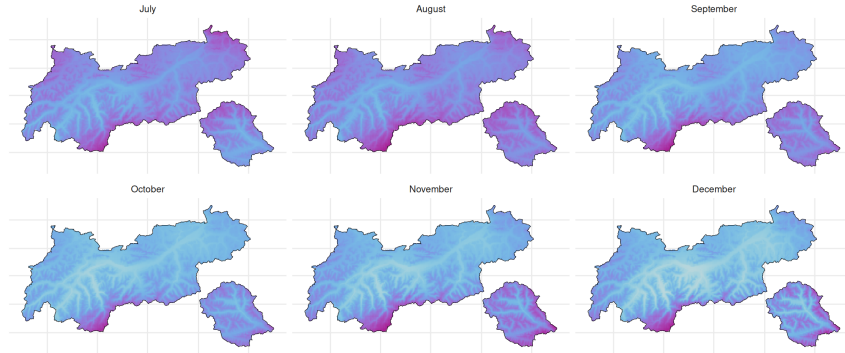
Application: Precipitation Climatology Estimation

Estimated climatology.



Application: Precipitation Climatology Estimation

Estimated climatology.



References

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