

Structural Equation Modeling for Social Relations: The R package `srm`

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Univariate Social Relations Model (I)

- actor i rates partner j in dyad $d = (ij)$ on one variable y , e.g., ratings on
 - *I like person XX a lot.*
 - *I think that person XX is good at Mathematics.*
- social relations model (SRM)

$$y_{ij} = \mu + a_i + p_j + \varepsilon_{ij} \quad (1)$$

- actor effects a_i : *how much person i likes other persons*
- partner effects p_j : *how much person j is liked by other persons*
- relationship effects ε_{ij} : *specific effect that person i likes j*

Univariate Social Relations Model (II)

- social relations model (SRM)

$$y_{ij} = \mu + a_i + p_j + \varepsilon_{ij} \quad (1)$$

- model parameters at level of persons (Σ_u) and dyads (Σ_r)

$$\Sigma_u = \text{Var} \begin{pmatrix} a_i \\ p_i \end{pmatrix} = \begin{pmatrix} \sigma_a^2 & \\ \sigma_{ap} & \sigma_p^2 \end{pmatrix} \quad (2)$$

$$\Sigma_r = \text{Var} \begin{pmatrix} \varepsilon_{ij} \\ \varepsilon_{ji} \end{pmatrix} = \begin{pmatrix} \sigma_\varepsilon^2 & \\ \sigma_{\varepsilon\varepsilon} & \sigma_\varepsilon^2 \end{pmatrix} \quad (3)$$

Mixed Effects Representation of the SRM

- social relations model (SRM)

$$y_{ij} = \mu + a_i + p_j + \varepsilon_{ij} \quad (1)$$

- define vector of person effects for persons $i = 1, \dots, I$: $\mathbf{u}_i = (a_i, p_i)$
- define vector of dyad effects for dyads $d = 1, \dots, D$: $\mathbf{r}_d = (\varepsilon_{ij}, \varepsilon_{ji})$
- collect all observations in outcome $\mathbf{y} = (y_{ij})_{ij}$
- mixed effects model representation (see Nestler, 2016)

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sum_{i=1}^I \mathbf{Z}_i \mathbf{u}_i + \sum_{d=1}^D \mathbf{W}_d \mathbf{r}_d \quad (4)$$

with design matrices \mathbf{Z}_i and \mathbf{W}_d (containing only zeros or ones)

- short form in mixed effects model notation: $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{W}\mathbf{r}$

Multivariate Social Relations Model

- now consider V multiple outcomes y_{1ij}, \dots, y_{Vij}
- multiple (i.e., $2V$) actor and partner effects define person level variable \mathbf{u}_i
- relationship vector \mathbf{r}_d can also be extended for multiple outcomes
- no general change in notation for mixed effects representation

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sum_{i=1}^I \mathbf{Z}_i \mathbf{u}_i + \sum_{d=1}^D \mathbf{W}_d \mathbf{r}_d \quad (4)$$

- in short: $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{W}\mathbf{r}$
- estimation with ANOVA method (unweighted least squares) or (restricted) maximum likelihood

Structural Equation Models (SEM) for Multivariate Data

- model multivariate normally distributed outcome as a constrained model $\mathbf{y} \sim MVN(\boldsymbol{\mu}(\boldsymbol{\theta}), \boldsymbol{\Sigma}(\boldsymbol{\theta}))$ with a parameter vector $\boldsymbol{\theta}$
- ignore mean structure in the following for simplicity
- structural equation model (SEM)

$$\begin{aligned} \mathbf{y} &= \boldsymbol{\Lambda}\boldsymbol{\eta} + \boldsymbol{\varepsilon} \\ \boldsymbol{\eta} &= \mathbf{B}\boldsymbol{\eta} + \boldsymbol{\xi} \end{aligned} \quad (5)$$

- model parameter vector $\boldsymbol{\theta}$ contains free parameters in $\boldsymbol{\Lambda}$, \mathbf{B} , $Var(\boldsymbol{\xi}) = \boldsymbol{\Phi}$, $Var(\boldsymbol{\varepsilon}) = \boldsymbol{\Psi}$
- model implied covariance matrix

$$Var(\mathbf{y}) = \boldsymbol{\Sigma}_y = \boldsymbol{\Sigma}_y(\boldsymbol{\theta}) = \boldsymbol{\Lambda}(\mathbf{I} - \mathbf{B})^{-1}\boldsymbol{\Phi}((\mathbf{I} - \mathbf{B})^{-1})'\boldsymbol{\Lambda}' + \boldsymbol{\Psi} \quad (6)$$

Maximum Likelihood Estimation in SEM

- maximum likelihood (ML) estimation maximizes

$$l(\boldsymbol{\theta}) = \text{const} - \frac{1}{2} \log |\boldsymbol{\Sigma}_{\mathbf{y}}^{-1}| - \frac{1}{2} (\mathbf{y} - \boldsymbol{\mu}_{\mathbf{y}})' \boldsymbol{\Sigma}_{\mathbf{y}}^{-1} (\mathbf{y} - \boldsymbol{\mu}_{\mathbf{y}}) \quad (7)$$

- gradient (score equation)

$$\frac{\partial l}{\partial \theta_h} = -\frac{1}{2} \text{tr} \left(\boldsymbol{\Sigma}_{\mathbf{y}}^{-1} \frac{\partial \boldsymbol{\Sigma}_{\mathbf{y}}}{\partial \theta_h} \right) + \frac{1}{2} (\mathbf{y} - \boldsymbol{\mu}_{\mathbf{y}})' \boldsymbol{\Sigma}_{\mathbf{y}}^{-1} \frac{\partial \boldsymbol{\Sigma}_{\mathbf{y}}}{\partial \theta_h} \boldsymbol{\Sigma}_{\mathbf{y}}^{-1} (\mathbf{y} - \boldsymbol{\mu}_{\mathbf{y}}) \quad (8)$$

- expected information matrix for use in Fisher Scoring

$$E \left(\frac{\partial^2 l}{\partial \theta_h \partial \theta_k} \right) = -\frac{1}{2} \text{tr} \left(\boldsymbol{\Sigma}_{\mathbf{y}}^{-1} \frac{\partial \boldsymbol{\Sigma}_{\mathbf{y}}}{\partial \theta_h} \boldsymbol{\Sigma}_{\mathbf{y}}^{-1} \frac{\partial \boldsymbol{\Sigma}_{\mathbf{y}}}{\partial \theta_k} \right) \quad (9)$$

- update equation in Fisher scoring

$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + \left[E \left(\frac{\partial^2 l}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \right) \right]^{-1} \frac{\partial l}{\partial \boldsymbol{\theta}} \quad (10)$$

Social Relations Structural Equation Model (SR-SEM)

- multivariate SRM \Rightarrow covariance structure of person effects Σ_u and dyad effects Σ_r
- consider restricted models $\Sigma_u = \Sigma_u(\theta)$ and $\Sigma_r = \Sigma_r(\theta)$, e.g. models with factor structures or relationship among several constructs \Rightarrow social relations structural equation model (SR-SEM)
- SEM at level of persons: $\theta_u = (\Lambda_u, B_u, \Phi_u, \Psi_u)$
- SEM at level of dyads $\theta_r = (\Lambda_r, B_r, \Phi_r, \Psi_r)$
- or pose some equality constraints among both levels (e.g., invariance of factor loadings)

ML Estimation in SR-SEM

- stack all observations (dyads, variables) of a round robin design in outcome vector \mathbf{y}
- \mathbf{y} is multivariate normally distributed if all effects of the SRM are normally distribution
- ML estimation of θ needs $\Sigma_{\mathbf{y}}$ and $\frac{\partial \Sigma_{\mathbf{y}}}{\partial \theta_h}$ (see normal theory based ML)
- multivariate SRM has mixed effects representation

$$\mathbf{y} = \mathbf{X}\beta + \sum_{i=1}^I \mathbf{Z}_i \mathbf{u}_i + \sum_{d=1}^D \mathbf{W}_d \mathbf{r}_d \quad (4)$$

⇒

$$\Sigma_{\mathbf{y}} = \text{Var}(\mathbf{y}) = \sum_{i=1}^I \mathbf{Z}_i \Sigma_u \mathbf{Z}_i' + \sum_{d=1}^D \mathbf{W}_d \Sigma_r \mathbf{W}_d' \quad (11)$$

$$\frac{\partial \Sigma_{\mathbf{y}}}{\partial \theta_h} = \sum_{i=1}^I \mathbf{Z}_i \frac{\partial \Sigma_u}{\partial \theta_h} \mathbf{Z}_i' + \sum_{d=1}^D \mathbf{W}_d \frac{\partial \Sigma_r}{\partial \theta_h} \mathbf{W}_d' \quad (12)$$

R Package `srm`

- R package `srm` on CRAN
- covers SEM at both levels (persons and dyads)
- satisfactory computation time (computational shortcuts, use of Rcpp)
- ML estimation using Fisher scoring and quasi-Newton approach using observed information matrix
- Fisher scoring relatively stable, at least more stable than Quasi-Newton algorithms with observed information matrix

srm Package: Model Syntax

- inspired by multilevel syntax of **lavaan** (level identifiers %person and %dyad)
- SRM decomposition $Y_{ij} = \mu + a_i + p_j + \varepsilon_{ij}$ plainly translates to $V1 = V1@A + V1@P + V1@AP$
- Example syntax for unidimensional factor model

```
\%Person
```

```
f1@A =~ Wert1@A + Wert2@A + Wert3@A
```

```
f1@P =~ Wert1@P + Wert2@P + Wert3@P
```

```
\%Dyad
```

```
f1@AP =~ Wert1@AP + Wert2@AP + Wert3@AP
```

```
# define some constraints
```

```
Wert1@AP ~~ 0*Wert1@PA
```

```
Wert3@AP ~~ 0*Wert3@PA
```

srm Package: Model Output

	index	group	lhs	op	rhs	mat	fixed	est	se	lower
1	NA	1	F1@A	=~	Wert1@A	LAM_U	1	1.000	NA	-Inf
2	NA	1	F1@P	=~	Wert1@P	LAM_U	1	1.000	NA	-Inf
3	1	1	F1@A	~~	F1@A	PHI_U	NA	0.322	0.071	-Inf
4	2	1	F1@A	~~	F1@P	PHI_U	NA	0.098	0.043	-Inf
5	3	1	F1@P	~~	F1@P	PHI_U	NA	0.160	0.049	-Inf
6	NA	1	Wert1@A	~~	Wert1@A	PSI_U	0	0.000	NA	-Inf
7	NA	1	Wert1@A	~~	Wert1@P	PSI_U	0	0.000	NA	-Inf
8	NA	1	Wert1@P	~~	Wert1@P	PSI_U	0	0.000	NA	-Inf
9	NA	1	F1@A	~1	F1@A	MU_U	0	0.000	NA	-Inf
10	NA	1	F1@P	~1	F1@P	MU_U	0	0.000	NA	-Inf
11	4	1	Wert1@A	~1	Wert1@A	BETA	NA	0.150	0.093	-Inf
12	NA	1	F1@AP	=~	Wert1@AP	LAM_D	1	1.000	NA	-Inf
13	NA	1	F1@PA	=~	Wert1@PA	LAM_D	1	1.000	NA	-Inf
14	6	1	F1@AP	~~	F1@AP	PHI_D	NA	1.531	0.081	-Inf
15	5	1	F1@AP	~~	F1@PA	PHI_D	NA	0.069	0.081	-Inf
16	6	1	F1@PA	~~	F1@PA	PHI_D	NA	1.531	0.081	-Inf
17	NA	1	Wert1@AP	~~	Wert1@AP	PSI_D	0	0.000	NA	-Inf
18	NA	1	Wert1@AP	~~	Wert1@PA	PSI_D	0	0.000	NA	-Inf
19	NA	1	Wert1@PA	~~	Wert1@PA	PSI_D	0	0.000	NA	-Inf

Computational Aspects

- matrices of derivatives $\frac{\partial \Sigma_u}{\partial \theta_h}$ and $\frac{\partial \Sigma_r}{\partial \theta_h}$ have known forms (known from single-level SEMs)
 - inverse matrix Σ_y^{-1} computationally demanding because its dimension is $D(D-1)V$
 - total likelihood based on sum of independent likelihoods corresponding to different round robin groups
- ⇒ Σ_y^{-1} must only be computed for round robin designs with same number of persons (without missing data)

Faster Computation of $\Sigma_{\mathbf{y}}^{-1}$: Woodbury Identity

- tip from Yves Rosseel (June 2019)
- observations in the SR-SEM are of the form $\mathbf{y} = \mathbf{Z}\mathbf{u} + \mathbf{e}$, where $\mathbf{U} = \text{Var}(\mathbf{u})$ and $\mathbf{E} = \text{Var}(\mathbf{e})$ are block diagonal matrices of functions of Σ_u and Σ_r , respectively
- Σ_u and Σ_r computationally inexpensive to invert (because of lower dimension), and, therefore, also block diagonal matrices \mathbf{U} and \mathbf{E}
- it holds that

$$\text{Var}(\mathbf{y}) = \Sigma_{\mathbf{y}} = \mathbf{Z}\mathbf{U}\mathbf{Z}^T + \mathbf{E} \quad (13)$$

- use Woodbury identity for inversion

$$(\mathbf{Z}\mathbf{U}\mathbf{Z}^T + \mathbf{E})^{-1} = \mathbf{E}^{-1} - \mathbf{E}^{-1}\mathbf{Z}(\mathbf{U}^{-1} + \mathbf{Z}^T\mathbf{E}^{-1}\mathbf{Z})\mathbf{Z}^T\mathbf{E}^{-1} \quad (14)$$

Skipping Zero Entries in Matrix Computations

- in computation of the first and second derivative, matrix multiplications $\Sigma_{\mathbf{y}}^{-1} \frac{\partial \Sigma_{\mathbf{y}}}{\partial \theta_h}$ for all parameters θ_h have to be computed
 - many entries in $\frac{\partial \Sigma_{\mathbf{y}}}{\partial \theta_h}$ are zero (e.g., derivative with respect to a particular item loading)
 - skip these computations in matrix computations by *hard coding sparse matrix multiplications* in **Rcpp**
- ⇒ skipping redundant computations led to most important speed improvement

More Advanced Models and Extensions

- multiple group models (e.g., round robin designs in different age groups or different school tracks)
- discrete moderators x (e.g., gender) of model parameters $\theta = \theta(x)$ can be handled by including pseudo variables (original variable \times dummy variables for moderator values)
- generic variables at person level (self ratings) are round robin variables with constraints: $y_{ij} = \mu + 0 \cdot a_i + 1 \cdot p_j + 0 \cdot \varepsilon_{ij}$
- level-specific fit indices for assessing differences between multivariate saturated SRM and SRM-SEM

Alternative Estimators

- least squares estimation (Bond & Malloy, 2018)
- composite likelihood methods (pairwise likelihood estimation), particularly attractive for high-dimensional models and categorical data
- MCMC techniques (Hoff, 2005; Gill & Swartz, 2001)
- maximum a posterior (MAP) estimation using prior distributions (penalized maximum likelihood estimation)
- plausible value imputation: estimate a saturated multivariate SRM at first, then plugin the PVs into a standard single-level SEM
- two-step methods: estimation of „factor scores“, then plug-in factor scores into path models (with some unreliability correction)

Many thanks!

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