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A Conceptional Lego Toolbox for Bayesian Distributional Regression Models

Existing Implementations:

► From univariate to multivariate

Highly specialized and optimized

► However, almost any engine has a

distributions.

different interface.

engines.

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1. INTRODUCTION

Bayesian Models:

- ► Convenient setting for very complex ► Provide infrastructures for a number of regression problems. models.
- MCMC simulation now possible on virtually any computer.
- Inference does not rely on asymptotics.

Basic Ideas:

- Provide a flexible and unified modeling architecture.
- ► Use specialized/optimized engines to apply Bayesian structured additive distributional regression a.k.a. Bayesian additive models for location scale and shape (BAMLSS) and beyond.
- ► Facilitate new algorithms and extensions.
- ► The approach should have maximum flexibility/extendability, also concerning functional types.

2. MODEL STRUCTURE

Within the GAMLSS model class all parameters of the response distribution can be modeled by explanatory variables

$$\mathbf{y} \sim \mathcal{D}\left(h_1(\theta_1) = \eta_1, \ h_2(\theta_2) = \eta_2, \ldots, \ h_K(\theta_K) = \eta_K\right)$$

where \mathcal{D} denotes any distribution available for the response variable y and θ_k , $k = 1, \ldots, K$, are parameters that are linked to additive predictors. The *k*-th additive predictor is given by

$$\eta_k = \eta_k(\mathbf{x}; \boldsymbol{\beta}_k) = f_{1k}(\mathbf{x}; \boldsymbol{\beta}_{1k}) + \ldots + f_{J_kk}(\mathbf{x}; \boldsymbol{\beta}_{J_kk}),$$

with unspecified (possibly nonlinear) functions $f_{jk}(\cdot)$ of a generic covariate vector **x**, $j = 1, ..., J_k$ and k = 1, ..., K. Examples of functions $f_{ik}(\cdot)$:



Terms:

Each vector of function evaluations is a composition of

 $\mathbf{f}_{jk} = f_{jk}(\mathbf{X}_{jk}, \boldsymbol{\beta}_{jk}),$

 \mathbf{X}_{jk} $(n \times m_{jk})$ is a design matrix, β_{jk} $(q_{jk} \times 1)$ are regression coefficients. Note: functions are not necessarily linear in the parameters, e.g., growth curves.

Priors:

Generic prior for linear and nonlinear effects using a basis function approach

$$p(\boldsymbol{eta}_{jk}) \propto \left(rac{1}{ au_{jk}^2}
ight)^{rk(\mathbf{K}_{jk})/2} \exp\left(-rac{1}{2 au_{jk}^2} \boldsymbol{eta}_{jk}^{ op} \mathbf{K}_{jk} \boldsymbol{eta}_{jk}
ight).$$

Precision matrix \mathbf{K}_{jk} corresponds to frequentist's penalty matrix, τ_{jk}^2 is equivalent to the inverse smoothing parameter, common prior $p(\tau_{jk}^2) \sim IG(a_{jk}, b_{jk})$.

Response Distribution:

Main building block pdf $f(\mathbf{y}|\boldsymbol{\theta}_1,\ldots,\boldsymbol{\theta}_K)$ and corresponding log-likelihood

$$\ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \log f(y_i; \theta_{i1} = h_1^{-1}(\eta_{i1}(\mathbf{x}_i, \boldsymbol{\beta}_1)), \dots, \theta_{iK} = h_K^{-1}(\eta_{iK}(\mathbf{x}_i, \boldsymbol{\beta}_K))).$$

Log-posterior

$$\log p(\boldsymbol{\vartheta} \cdot \boldsymbol{\mathbf{x}} \boldsymbol{\mathbf{X}}) \propto \ell(\boldsymbol{\vartheta} \cdot \boldsymbol{\mathbf{x}} \boldsymbol{\mathbf{X}}) + \sum_{k=1}^{K} \sum_{j=1}^{J_k} \{\log p_{ij}(\boldsymbol{\vartheta}_{ij})\}$$

An implementation is provided in the R package **bamlss** available at

https://R-Forge.R-project.org/projects/BayesR/

In R, simply type

> install.packages("bamlss", repos = "http://R-Forge.R-project.org")

Generic architecture, the setup does not restrict to any specific type of engine (Bayesian or frequentist).

Various algorithms implemented, in addition support for **BayesX**, JAGS, Stan.



5. EXAMPLE

Modeling daily precipitation data with a censored normal model $\mathbf{y}^{\star} \sim \mathcal{N}(oldsymbol{\mu}, oldsymbol{\sigma}^2), \quad oldsymbol{\mu} = oldsymbol{\eta}_{\mu}, \quad \log(oldsymbol{\sigma}) = oldsymbol{\eta}_{\sigma}, \quad \mathbf{y} = max(\mathbf{0}, \mathbf{y}^{\star}).$ For both μ and σ , we use the following additive predictor: $\eta = \beta_0 + f_1(\text{day}, \text{lon}, \text{lat}) + f_2(\text{lon}, \text{lat}) + f_3(\text{day}) + f_4(\text{alt}).$

 $\log p(\boldsymbol{v}, \boldsymbol{y}, \boldsymbol{\Lambda}) \propto \ell(\boldsymbol{\rho}, \boldsymbol{y}, \boldsymbol{\Lambda}) + \sum \{\log p_{jk}(\boldsymbol{v}_{jk})\},\$ k=1 i=1

e.g., $\vartheta_{jk} = (\beta_{jk}^{\top}, (\tau_{jk}^2)^{\top})^{\top}$ and $p_{jk}(\cdot)$ denotes combination of all priors. **Posterior Mode/Mean Estimation:**

Various algorithms require

 $\blacktriangleright \frac{\partial \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\beta}_{k}} = \frac{\partial \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\eta}_{k}} \frac{\partial \boldsymbol{\eta}_{k}}{\partial \boldsymbol{\beta}_{k}} = \frac{\partial \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\theta}_{k}} \frac{\partial \boldsymbol{\theta}_{k}}{\partial \boldsymbol{\eta}_{k}} \frac{\partial \boldsymbol{\eta}_{k}}{\partial \boldsymbol{\beta}_{k}},$ $\bullet \frac{\partial^2 \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\beta}_k \partial \boldsymbol{\beta}_s^{\top}} = \left(\frac{\partial \boldsymbol{\eta}_s}{\partial \boldsymbol{\beta}_s}\right)^{\top} \frac{\partial^2 \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\eta}_k \partial \boldsymbol{\eta}_s^{\top}} \frac{\partial \boldsymbol{\eta}_k}{\partial \boldsymbol{\beta}_k} + \frac{\partial \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\eta}_k} \frac{\partial^2 \boldsymbol{\eta}_k}{\partial^2 \boldsymbol{\beta}_k} \text{ if } k = s,$ $\blacktriangleright \frac{\partial^2 \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\eta}_k \partial \boldsymbol{\eta}_s^{\top}} = \frac{\partial \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\theta}_k} \frac{\partial^2 \boldsymbol{\theta}_k}{\partial \boldsymbol{\eta}_k \partial \boldsymbol{\eta}_s^{\top}} + \frac{\partial^2 \ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X})}{\partial \boldsymbol{\theta}_k \partial \boldsymbol{\theta}_s^{\top}} \frac{\partial \boldsymbol{\theta}_k}{\partial \boldsymbol{\eta}_k} \frac{\partial \boldsymbol{\theta}_s}{\partial \boldsymbol{\eta}_s}.$

Hence, implementing new distributions usually requires derivatives for η_k , only. Generic (blockwise) iterative updating scheme

 $\boldsymbol{\beta}_{k}^{(t+1)} = U_{k}(\boldsymbol{\beta}_{k}^{(t)}|\cdot),$

e.g., for backfitting or MCMC.



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